

RUDOLF CARNAP

NOTES ON PROBABILITY AND INDUCTION*

INTRODUCTORY REMARKS

I. *The Three Main Conceptions of Probability*

- (1) The classical conception (Bernoulli, Bayes, Laplace).
- (2) The frequency conception (Mises, Reichenbach; mathematical statistics).
- (3) The logical conception (Keynes, Jeffreys).

Read: Nagel [37]

II. *The Two Explicanda*

There are two explicanda, both called 'probability':

- (1) logical or inductive probability (probability₁),
- (2) statistical probability (probability₂).

Read: [Prob.] Ch. II, esp. §§ 9 and 10.

The logical concept of probability appears in three forms ([Prob.] § 8):

- (a) the classificatory concept (confirming evidence),
- (b) the comparative concept (higher confirmation),
- (c) the quantitative concept (degree of confirmation).

III. *Preliminary Remarks on Inductive Logic*

Read: [Prob.] Ch. IV. In particular:

- (1) Logical probability (as explicandum) is explained as a *fair betting quotient*, and as an estimate of relative frequency ([Prob.] § 41).
- (2) If logical probability is used, no synthetic assumption (e.g., *uniformity* of the world) is needed as presupposition for the validity of the inductive method ([Prob.] § 41 F).
- (3) Comparison of *inductive and deductive logic* ([Prob.] § 43).
- (4) The main kinds of *inductive inference* ([Prob.] § 44 B)
 - (a) direct inference (from the population to a sample),

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- (b) predictive inference (from one sample to another),
- (c) inference by analogy,
- (d) inverse inference (from a sample to the population),
- (e) universal inference (from a sample to a universal law).

(5) The use of inductive logic for the choice of a *practical decision* ([Prob.] §§ 50, 51). The rule of maximizing the estimate of utility. Daniel Bernoulli's law of utility.

IV. *Some Concepts of Deductive Logic*

Read: [Prob.] §§ 14-20, esp. 18-20.

(1) *State-descriptions* (Z, § 18A; comp. individual distributions, D 26-6a). A state-description describes a (possible) state or *model*.

(2) The requirement of the logical independence of the primitive predicates (§ 18B) can be abandoned if the dependences are expressed by *meaning postulates* (see [15].)

(3) *Families of related primitive predicates* (§ 18C).

(4) The *range* of a sentence (§§ 18D, 19).

(5). *L-concepts* (§ 20). I write ' $\vdash i$ ' for '*i* is L-true', and hence ' $\vdash i \supset j$ ' for '*i* L-implies *j*', and ' $\vdash i \equiv j$ ' for '*i* is L-equivalent to *j*'.

In simple languages (e.g., those used in [Prob.]) every model is describable by a state-description. In richer languages this is not possible; here the definitions of L-concepts and degree of confirmation are to be based on *models* rather than state-descriptions.

THE THEORY OF DEGREE OF CONFIRMATION

V. *Fundamental axioms* (A1-A5)

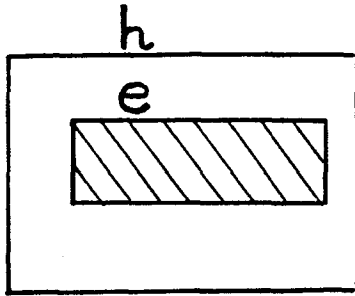
The axioms apply to any sentences *e* and *h* in a given language *L* (finite; or infinite). We presuppose throughout that the second argument of *c* (usually *e*) is not L-false (see [Prob.] pp. 295f.).

- A1.** *Range of values.* $0 \leq c(h, e) \leq 1$.
- A2.** *L-implication.* If $\vdash e \supset h$, then $c(h, e) = 1$.
- A3.** *Special addition principle.* If $e \cdot h \cdot h'$ is L-false, then $c(h \vee h', e) = c(h, e) + c(h', e)$.
- A4.** *General multiplication principle.* $c(h, h', e) = c(h, e) \times c(h', e \cdot h)$.

A5. *L-equivalent arguments.* If $\vdash e \equiv e'$ and $\vdash h \equiv h'$, then $c(h, e) = c(h', e')$.

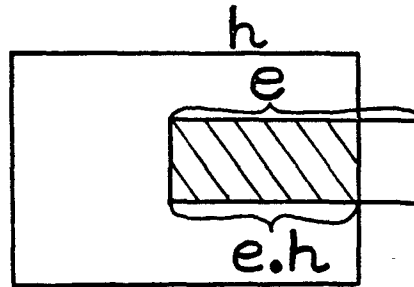
(These axioms, except for A5, are those of Shimony [47]. They are together equivalent to the Conventions C53-1 and 2 in [Prob.] § 53. Most axiom systems of other authors are essentially equivalent to this one; see [Prob.] § 62.) The usual theorems of the probability calculus are provable on the basis of these axioms. Among them are the theorems [Prob.] T53-1a to f. ('*t*' is the tautology.)

VI. *Regular m-Functions and c-Functions* (A6)



Deductive logic

'*e* L-implies *h*' means that the range of *e* is entirely contained in that of *h*.



Inductive logic

' $c(h, e) = 3/4$ ' means that three-fourths of the range of *e* is contained in that of *h*. ([Prob.] § 55B.)

Fig. 1.

For degrees of confirmation (d. of c.) we need a measure function for the ranges of sentences. For this purpose we define regular *m*-functions ([Prob.] 55A).

D1. *m* is a regular *m*-function for $L_N =_{\text{Df}}$

- (a) for every Z_i in L_N , $m(Z_i) > 0$;
- (b) $\sum_i m(Z_i) = 1$;
- (c) if *j* is L-false, $m(j) = 0$;
- (d) if *j* is not L-false, $m(j) = \sum m(Z_i)$ for all Z_i in the range of *j*.

D2. *c* is a regular *c*-function for $L_N =_{\text{Df}}$ there is a regular *m*-function *m* such that *c* is based upon *m*, i.e.

$$c(h, e) = \frac{m(e.h)}{m(e)}.$$

A6. *Regularity.* In a finite domain of individuals, $c(h, e) = 1$ only if $\vdash e \supset h$.

(This axiom corresponds to [Prob.] C53-3.)

Null confirmation is the d. of c . on the tautological evidence t ([Prob.] D57-1, where the symbol ' c_0 ' is used):

D3. $c_t(j) =_{\text{Df}} c(j, t)$

T1. A c -function c for L_N satisfies the axioms A1-A6 if and only if c is a regular c -function.

Proof. 1. Let c be a regular c -function for L_N . Then c satisfies A1-A6 according to [Prob.] T59-1a, 1b, 1l, 1n, 1h and i, T59-5a, respectively. 2. Let c satisfy A1-A6. Then c_t is a regular m -function (by [Prob.] C53-3 and T53-1). c is based upon c_t (comp. [Prob.] § 54B, (3)). Therefore c is a regular c -function.

According to T1, the theorems stated in [Prob.] §§ 55, 57A and B, 59, 60, and 61 for regular c -functions in finite systems L_N are provable on the basis of A1 to A6.

If c satisfies A1 to A5, but not A6, we shall call it a quasi-regular c -function (not in [Prob. j]). In this case, c_t is 0 for some Z_i ; therefore, even in L_N , $c(h, e)$ cannot always be represented as $c_t(e \cdot h)/c_t(e)$.

(Example: *the straight rule*, [Prob.] p. 227.)

The following theorems are provable on axioms A1 to A5; hence they hold for all regular or quasi-regular c -functions.

T2. $c(h, e) \times c(i, e \cdot h) = c(i, e) \times c(h, e \cdot i)$. (From A4, A5.)

T3. General division theorem, in two forms.

a. If $c(i, e) > 0$, then $c(h, e \cdot i) = \frac{c(h, e) \times c(i, e \cdot h)}{c(i, e)}$ (From T2.)

b. If $c(i, e) > 0$ and $c(h, e) > 0$, then $\frac{c(h, e \cdot i)}{c(h, e)} = \frac{c(i, e \cdot h)}{c(i, e)}$ (From (a).)

The fraction on the left-hand side of the equation is known as the

relevance quotient; the numerator is the *posterior confirmation* of h and the denominator is the *prior confirmation* of h .

T4. *Special division theorem.* Suppose that $c(i, e) > 0$, $c(h, e) > 0$, and $c(i, e \cdot h) = 1$ (i is predictable or explainable by h). Then

$$\frac{c(h, e \cdot i)}{c(h, e)} = \frac{1}{c(i, e)} \text{ (From T3b).}$$

See the explanations and examples for these theorems in [Prob.] §§ 60 and 61.

VII. Coherence

Informal explanation. Let X be willing to accept any system of bets in which the betting quotients are equal to the values of a function c . If there were a betting system such that X would suffer a loss in every logically possible case, c would obviously be unsuitable. If there is no such betting system, we shall call c *coherent* (Ramsey, De Finetti). If, moreover, there is no betting system such that X would lose in at least on possible case and would not gain in any possible case, we shall call c *strictly coherent* (Shimony).

We assume for the following definitions that \mathbb{L} is an interpreted language, that e and h are sentences of \mathbb{L} , that e is not \mathbb{L} -false, that c is a function whose value for any h, e is a real number, and that q and S are real numbers (and likewise for e_i, h_i, q_i, S_i).

We represent a bet (of the person X) on h , given e , in language \mathbb{L} , with the betting quotient q and the total stake S as the ordered quintuple $\langle \mathbb{L}, h, e, q, S \rangle$ (without reference to X):

D1. B is a *bet* =_{df} for some \mathbb{L}, h, e, q , and S , $B = \langle \mathbb{L}, h, e, q, S \rangle$.

We represent a betting system BS based on the assumption k and comprising the bets B_1, B_2, \dots, B_n in language \mathbb{L} , in accordance with c (i.e., the betting quotients are determined by the values of c) as the ordered quadruple $\langle \{B_1, \dots, B_n\}, \mathbb{L}, k, c \rangle$:

D2. BS is a *betting system* =_{df} $BS = \langle K, \mathbb{L}, k, c \rangle$, where $K = \{B_i\}$ ($i = 1, \dots, n$), $B_i = \langle \mathbb{L}, h_i, e_i, q_i, S_i \rangle$, k is a non- \mathbb{L} -false sentence in \mathbb{L} , each e_i is either k or a conjunction containing k as a component and is not \mathbb{L} -false, and $q_i = c(h_i, e_i)$.

If X regards a bet on h , given e , with betting quotient q as fair, then he is willing to make a corresponding bet on either side, i.e., either for h or against h . If e is true, the gains are as follows (with $S > 0$):

		Gain	
		h	
		for h	against h
(a)	true	$(1 - q)S$	$-(1 - q)S$
(b)	false	$-qS$	qS

Thus X 's bet against h can be regarded as a bet for h with negative S . Therefore we admit $S \leq \geq 0$; then D3 covers both bets, for h and against h . $g(B, j)$ is the gain which X would obtain from his bet B if j were true.

D3. Let B be a bet $\langle \mathbb{L}, h, e, q, S \rangle$. Let j be a non- \mathbb{L} -false sentence in \mathbb{L} which \mathbb{L} -implies either e or $\sim e$ and \mathbb{L} -implies either h or $\sim h$. $g(B, j) =_{\text{Df}}$ the value u such that

- either (a) $\vdash j \supset e \cdot h$, and $u = (1 - q)S$,
or (b) $\vdash j \supset e \cdot \sim h$, and $u = -qS$,
or (c) $\vdash j \supset \sim e$, and $u = 0$.

We define $G(\text{BS}, j)$ as the total gain from the betting system BS which X would obtain if j were true:

D4. Let BS be $\langle \{B_i\}, \mathbb{L}, k, c \rangle$ ($i = 1, \dots, n$). Let j be a non- \mathbb{L} -false sentence in \mathbb{L} such that, for every i , j \mathbb{L} -implies either e_i or $\sim e_i$, and j \mathbb{L} -implies either h_i or $\sim h_i$. Then $G(\text{BS}, j) =_{\text{Df}} \sum_{i=1}^n g(B_i, j)$.

Let BS be $\langle \{B_i\}, \mathbb{L}, k, c \rangle$. Let C_{BS} be the class of the conjunctions j such that (1) j contains as components, for each of the sentences $e_1, \dots, e_n, h_1, \dots, h_n$, either the sentence itself or its negation but not both, and no other components, and (2) j is compatible with k . These conjunction represent the possible cases on the basis of the assumption k . We shall say that for a given BS *loss is necessary* if, for every conjunction j in C_{BS} , $G(\text{BS}, j) < 0$; that *loss is possible* if, for at least one j in C_{BS} , $G(\text{BS}, j) < 0$; and that *positive gain is impossible* if, for every j in C_{BS} $G(\text{BS}, j) \leq 0$. We shall say that BS is *vacuous* if, for every j , $G(\text{BS}, j) = 0$.

- D5.** c is a *coherent c-function* for $\mathbb{L} =_{\text{Df}}$ there is no betting system in \mathbb{L} in accordance with c for which loss is necessary (in other words, for every betting system there is a possible outcome without loss).
- D6.** c is a *strictly coherent c-function* for $\mathbb{L} =_{\text{Df}}$ there is no betting system in \mathbb{L} in accordance with c for which loss is possible and positive gain is impossible (in other words, for every non-vacuous betting system there is a possible outcome with positive gain).
- T1.** If c is strictly coherent, it is also coherent.
- T2.** (*Ramsey, De Finetti.*) Every *coherent c-function* satisfies the axioms A1 to A5. In other words, if c violates at least one of the axioms A1 to A5, then there is betting system in accordance with c for which loss is necessary.

Example for A4. Suppose that c violates A4 in \mathbb{L} . Then there are sentences e , h , and h' in \mathbb{L} such that

$$c(h, e) \times c(h', e \cdot h) - c(h \cdot h', e) \neq 0.$$

Let $c_1 = c(h, e)$, $c_2 = c(h', e \cdot h)$, $c_3 = c(h \cdot h', e)$, and let $c_1 c_2 - c_3 = D$.

We choose the betting system $\text{BS} = \langle \{B_i\}, \mathbb{L}, e, c \rangle$ ($i = 1, 2, 3$),

TABLE I
Example for A4

	B_i	$g(B_i, j)$ for the four conjunctions in C_{BS}					
i	e_i	h_i	S_i	q_i	$e \cdot h \cdot h'$	$e \cdot h \cdot \sim h'$	$e \cdot \sim h \cdot h'$ $e \cdot \sim h \cdot \sim h'$
1	e	h	$\frac{c_2}{D}$	c_1	$\frac{(1-c_1)c_2}{D}$	$\frac{(1-c_1)c_2}{D}$	$\frac{-c_1 c_2}{D}$
2	$e \cdot h$	h'	$\frac{1}{D}$	c_2	$\frac{1-c_2}{D}$	$\frac{-c_2}{D}$	0
3	e	$h \cdot h'$	$\frac{-1}{D}$	c_3	$\frac{-(1-c_3)}{D}$	$\frac{c_3}{D}$	$\frac{c_3}{D}$
$G(\text{BS}, j) =$					-1	-1	-1

with e as k , and with e_i, h_i , and S_i as specified in the table below. (See Table I.) By D2, $q_i = c(h_i, e_i)$. The values of g are determined by D3, and those of G by D4. We find that, for every j , $G = -1$. Thus for the chosen BS, loss is necessary. This betting system is described in Table I.

T3. (*Shimony.*) If c violates A6, then there is a betting system in accordance with c for which loss is possible and positive gain is impossible. Therefore every *strictly coherent* c -function satisfies the axioms A1 to A6.

Proof. Suppose that c violates A6 in L_N . Then there are sentences e, h in L_N such that $c(h, e) = 1$ but e does not L-imply h , hence $e \cdot \sim h$ is not L-false. We take a system of one bet $\langle L_N, h, e, c(h, e), 1 \rangle$, and again e as k . The two possible cases j are $e \cdot h$ and $e \cdot \sim h$. The gain is 0 in the first case, and -1 in the second. Thus loss is possible and positive gain is impossible. This applies to any quasi-regular c -function, e.g. to the *straight rule* (VI).

T2 gives a *validation for the axioms A1 to A5, T3 for A6*. The following theorem shows that an analogous validation is not possible for any further axioms. (The proof for T4 is given by Kemeny in his paper [34].)

T4. (Kemeny) **a.** Every c -function in L which satisfies the axioms A1 to A5, is *coherent* in L .
 b. Every c -function in L which satisfies the axioms A1 to A6, is *strictly coherent* in L .

T5. **a.** A c -function is coherent if and only if it is regular or quasi-regular.
 b. A c -function is strictly coherent if and only if it is regular. (From T2, T3, T4, and VI-T1.)

The classification of c -functions defined by T4 and T5 can be presented in the form of the following table:

Axioms satisfied		Type of c -function	
A1 to A5	A6	regular strictly coherent	coherent
	not A6	quasi-regular	

VIII. Symmetrical c -Functions (A7)

The system A1 to A6 is very weak. It determines no value of $c(h, e)$ except 0 or 1 in special cases. For any pair of factual sentences e, h such that e L-implies neither h nor $\sim h$, the system does not exclude any number between 0 and 1 as a value of $c(h, e)$ ([Prob.] T59-5f, see remark on p. 323). Thus additional axioms are needed. A7 is the first of several *axioms of invariance* of $c(h, e)$ with respect to certain transformations of e and h . These axioms represent the valid core of the classical *principle of indifference*. *Axiom of symmetry* (with respect to individuals):

A7. $c(h, e)$ is invariant with respect to any permutation of the individuals.

DI. m -functions and c -functions which satisfy A7 are said to be *symmetrical* (with respect to individuals). (See [Prob.] §§ 90, 91.)

Read the definitions and explanations of the following concepts in [Prob.] : Ch. III: *division* (D25-4), *isomorphic* sentences (D26-3) and *isomorphic state-descriptions* (§ 27), *individual* and *statistical distributions* (D26-6), *structures* (§ 27) and *structure-descriptions* (*Str*, D27-1), *Q-predicates* (§ 31) and *Q-numbers* (§ 34).

Henceforth it is assumed, unless the contrary is stated, that c satisfies A1 to A7 and hence is regular and symmetrical. m is c_i ; hence c is based on m .

T1. Let e be isomorphic to e' , and h to h' .

a. $c(h, e) = c(h', e')$. (From A7).

b. $m(h) = m(h')$. (From (a)).

T2. Let i be an individual distribution for n given individuals with respect to the division M_1, \dots, M_k , with the cardinal numbers n_1, \dots, n_k .

a. The numbers of the individual distributions for the same n individuals which are isomorphic to i is

$$\zeta_i = \frac{n!}{n_1! \dots n_k!} \text{ ([Prob.] T40-32b.)}$$

b. Let j be the statistical distribution corresponding to i . Then $m(j) = \zeta_i \times m(i)$.
(From T1b).

T3 is a special case of T2.

T3. Let L_N be a language with N individual constants and k Q -predicates. Let Z_i be a state-description in L_N with the Q -numbers N_j ($j = 1, \dots, k$).
a. The number of those state-descriptions in L_N which are isomorphic to Z_i is

$$\zeta_i = \frac{N!}{N_1! N_2! \dots N_k!}. \text{ (From T2a.)}$$

b. Let Str_i be the structure-description corresponding to Z_i . Then $m(Str_i) = \zeta_i \times m(Z_i)$. (From T2b.)

Therefore a regular and symmetrical m -function for L_N is uniquely determined if we choose as its values for the structure-descriptions in L_N arbitrary positive numbers whose sum is 1. Then, for any Z_i , $m(Z_i)$ is determined by T3b and hence the other values by VI-D1c and d.

The subsequent *theorems* T4 to T6 on the direct inductive inference refer to the following situation. e is a statistical distribution for n given individuals (the ‘population’) in L_N with respect to the division M_1, M_2 (which is non- M_1) with the cardinal numbers n_1, n_2 . $r_i = n_i/n$ ($i=1, 2$). h is an individual distribution for s of the n individuals (the ‘sample’) with the cardinal numbers s_1, s_2 ($s_i \leq n_i$). h_{st} is the statistical distribution corresponding to h .

$$\mathbf{T4.} \quad \mathbf{a.} \quad c(h, e) = \frac{\begin{bmatrix} n_1 \\ s_1 \end{bmatrix} \begin{bmatrix} n_2 \\ s_2 \end{bmatrix}}{\begin{bmatrix} n \\ s \end{bmatrix}}.$$

(For $\begin{bmatrix} n \\ s \end{bmatrix}$, see [Prob.] D40-3.)

$$\mathbf{b.} \quad c(h_{st}, e) = \frac{\binom{n_1}{s_1} \binom{n_2}{s_2}}{\binom{n}{s}}$$

(For $\binom{n}{m}$, see D40-2.)

c. For given e and s , $c(h_{st}, e)$ has its maximum if s_1/s is equal, or as near as possible, to r_1 .

d. For fixed s , let $h_p(p = 0, \dots, s)$ be the statistical distribution h_{st} with $s_1 = p$ and $s_2 = s - p$. Then

$$\sum_{p=0}^s [p \times c(h_p, e)] = sr_1.$$

e. Let j be a full sentence of 'N₁' with one of the n individual constants in e . Then

$$c(j, e) = r_1. \quad (\text{For proofs see [Prob.] T94-1.})$$

We see from T4d that, for given s , the estimate of s_1 on e is sr_1 . Hence the estimate of s_1/s is r_1 . T4e shows that c for a singular prediction with 'N₁' is r_1 . Thus for the direct inference something analogous to the straight rule holds for all symmetrical regular (or quasi-regular) c -functions.

T5. The following holds *approximately* for sufficiently large n , n_1 , and n_2 . It holds exactly for $\lim c(n \rightarrow \infty)$ if $\lim(n_i/n) = r_i$.

$$\mathbf{a.} \quad c(h, e) = r_1^{s_1} \times r_2^{s_2}.$$

$$\mathbf{b.} \quad \text{Binomial law. } c(h_{st}, e) = \binom{s}{s_1} r_1^{s_1} r_2^{s_2}.$$

For proofs and explanations, see [Prob.] § 95.

We shall use the following notations in T6: $\sigma = \sqrt{sr_1r_2}$ ('standard deviation'); $\delta = s_1 - sr_1$ (deviation of s_1 from its estimate); $\phi(u) = (1/\sqrt{2\pi})e^{-u^2/2}$ (the normal function; [Prob.] D40-4a); h_p as in T4d; h' is the disjunction of sentences h_p with p running from $sr_1 - \delta'$ (or the integer nearest to it) to $sr_1 + \delta'$ ($= s_1'$); thus h' says that s_1 deviates from its estimate sr_1 to either side by not more than δ' , in other words, that s_1/s (the relative frequency of M_1 in the sample) does not deviate from r_1 by more than δ'/s .

T6. The following holds *approximately* for sufficiently large s and n/s .

a. *The normal law.*

$$c(h_{st}, e) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\delta^2/2\sigma^2} = \frac{1}{\sigma} \phi\left(\frac{\delta}{\sigma}\right).$$

b. *Bernoulli-Laplace theorem*

$$c(h', e) = \int_{-\delta'/\sigma}^{+\delta'/\sigma} \phi(u) du.$$

c. *Bernoulli's limit theorem.* For fixed r_1 and fixed $q = \delta'/s$, $\lim_{s \rightarrow \infty} c(h', e) = 1$.

T6c says the following. If the sample size s increases but a fixed interval $r_1 \pm q$ around the given r_1 is chosen, then $c(h', e)$ (i.e., the probability that the relative frequency of M_1 in the sample lies within the chosen interval) can be brought as near to 1 as desired by making the sample sufficiently large. For explanations and numerical examples, see [Prob.] § 96.

IX. Estimation

Read: [Prob.] § 98 about the present situation of the problem of estimation.

Definition of the general estimate function.

Suppose that, on the basis of e , the magnitude u has n possible values: u_1, \dots, u_n . Let h_i say that u has the value u_i ($i = 1, \dots, n$). The *c-mean estimate* of u is the weighted mean of the possible values, with their *c-values* as weights:

D1. $est(u, e) = \text{Df} \sum_{i=1}^n [u_i \times c(h_i, e)].$

T1. A and B are arbitrary fixed constants.

a. $est(Au, e) = A \times est(u, e).$

b. $est(u + B, e) = est(u, e) + B.$

c. $est(Au + B, e) = A \times est(u, e) + B.$

([Prob.] T100-3,4, and 5).

Analogous results do not generally hold for a non-linear function of u . For example, in general $est(u^2, e) \neq est^2(u, e)$. This leads to a *paradox*

in the practical application of estimates ([Prob.] § 100 C). The paradox is eliminated if the rule for the determination of a decision refers to the estimate of only one magnitude, e.g., the gain or the utility resulting from an action.

Truth frequency. Let K be a class of s sentences i_1, \dots, i_s . Let $tf(K)$ be the truth-frequency in K , i.e., the number of true sentences in K . Let $rtf(K)$ be the relative truth-frequency in K , i.e., $tf(K)/s$.

T2. **a.** $est(tf, K, e) = \sum_{n=1}^s c(i_n, e)$. (For this proof. see [Prob.] T104-2a.)

b. $est(rtf, K, e) = \frac{1}{s} \sum_{n=1}^s c(i_n, e)$. (From (a). T1a.)

c. If all sentences in K have the same c -value on e , then the estimate of $rtf(K)$ is equal to this c -value. (From (b).)

The frequency of a property of individuals. Let K be a class of n individuals defined by enumeration. Let $af(M, K)$ be the absolute frequency of M in K , and $rf(M, K)$ the relative frequency, i.e., of $(M, K)/n$. Let K' be the class of the full sentences of M with those individual constants which designate the individuals in K Then

$$af(M, K) = tf(K') \quad \text{and} \quad rf(M, K) = rtf(K').$$

Therefore the results T2 on estimates of tf and rtf can now be applied to estimates of af and rf .

Direct estimation of frequency. This is based on the direct inference (see VIII-T4). Let e , n , M_1 , n_1 , r_1 , s , and s_1 be as before (VIII-T4). Thus e says that the rf of M_1 in the population is r_1 . Let K be the class of the s individuals of the sample.

T3. **a.** $est(af, M, K, e) = sr_1$. (From VIII-T4d.)

b. $est(rf, M, K, e) = r_1$. (From (a), T1a.)

Predictive estimation of frequency. Here the estimate depends on the chosen c -function. Let e be any non-L-false sentence, h a full sentence of M for a new individual, and K any finite, non-empty class of new individuals.

T4. $est(rf, M, K, e) = c(h, e)$. (From T2c.)

Thus the confirmation of a singular prediction with M is equal to the estimate of rf of M . This relation was used earlier for an informal explanation of inductive probability ([Prob.] § 41D).

X. *The Functions e^\dagger and c^**

In discussions on the principle of indifference, some authors have proposed to give equal *a priori* probabilities to all individual distributions (for a given domain of individuals and a given division of properties). Other authors have proposed the same for all statistical distributions. In our terminology, the controversy concerns the choice of one of the following two rules:

- (A) All individual distributions have equal m -values.
- (B) All statistical distributions have equal m -values.

However, each of these rules leads to contradictions if applied to different divisions (see the examples in [Continuum] p. 39).

Each of the rules becomes consistent if it is restricted to one division (for a given finite language), viz. the division of the Q -predicates, as follows:

- (A') All state-descriptions have equal m -values.
- (B') All structure-descriptions have equal m -values.

The function c^\dagger . There is exactly one symmetrical, regular m -function which fulfills (A'), viz. m^\dagger defined by D1.

Let \mathbb{L}_N be a language with N individual constants and k Q -predicates.

- T1.** **a.** The number of state-descriptions in \mathbb{L}_N is $\zeta_N = k^N$. ([Prob.] T40-31c.)
 b. The number of structure-descriptions in \mathbb{L}_N is

$$\tau_N = \binom{N+k-1}{k-1} = \frac{(N+k-1)!}{N!(k-1)!}. \quad ([\text{Prob.}] \text{ T40-33b.})$$

Let Z_N be any state-description in \mathbb{L}_N with the Q -numbers N_1, \dots, N_k . We define:

D1. $m^\dagger(Z_N) =_{\text{Df}} \frac{1}{k^N}$

$$\mathbf{D2.} \quad c^\dagger(h, e) =_{\text{Df}} \frac{m^\dagger(e.h)}{m^\dagger(e)}.$$

Let e_N be an individual distribution for any N individuals for the division of the k Q -predicates with the same Q -numbers N_1, \dots, N_k (the same as in Z_N). Let h_j be a full sentence of Q_j for a new individual.

T2. a. m^\dagger is regular and symmetrical. (From D1.)

$$\mathbf{b.} \quad m^\dagger(e_N) = \frac{1}{k^N}. \quad (\text{From D1, since } e_N \text{ is isomorphic to } Z_N.)$$

$$\mathbf{c.} \quad c^\dagger(h_j, e_N) = 1/k.$$

Proof. $e_N \cdot h_j$ is isomorphic to a state-description in L_{N+1} , hence $m^\dagger = 1/k^{N+1}$ (from D1). The result is obtained by D2 and (b).

T2c shows that $c^\dagger(h_j, e_N)$ is independent of e_N . It violates the principle of learning from experience and hence is unacceptable ([Prob.] p. 565). However, this function was proposed by C. S. Peirce, Keynes, and Wittgenstein.

The function c^ .* There is exactly one symmetrical, regular m -function which fulfills (B'), viz. m^* defined by D3.

$$\begin{aligned} \mathbf{D3.} \quad m^*(Z_N) &=_{\text{Df}} \frac{1}{\tau_N \zeta_i} \\ &= \frac{N_1! \dots N_k! (k-1)!}{(N+k-1)}. \quad (\text{From T1b, VIII-T3a.}) \end{aligned}$$

T3. a. For any structure-description in L_N , $m^* = \frac{1}{\tau_N}$. Thus m^* fulfills (B). (From VIII-T3b.)

b. m^* is regular and symmetrical. (From D3.)

$$\mathbf{c.} \quad m^*(e_N) = \frac{N_1! \dots N_k! (k-1)!}{(N+k-1)}. \quad (\text{From D3.})$$

c^* is based on m^* :

$$\mathbf{D4.} \quad c^*(h, e) =_{\text{Df}} \frac{m^*(e.h)}{m^*(e)}.$$

$$\mathbf{T4.} \quad c^*(h_j, e_N) = \frac{N_j + 1}{N + k}.$$

Proof. $e_N \cdot h_j$ is isomorphic to a state-description in L_{N+1} with the Q -numbers $N_1, \dots, N_j + 1, \dots, N_k$. Therefore its m^* -value is like that of e_N in T3c, but with $N_j + 1$ instead of N_j and $N + 1$ instead of N . Hence the result by D4.

Let M be a disjunction of w Q -predicates ($0 < w < k$) and N_M be the sum of the Q -numbers of these Q -predicates in e_N . Hence w is the *logical width* of M ([Prob.] § 32). Let h_M be a full sentence of M for a new individual.

$$\mathbf{T5.} \quad c(h_M, e_N) = \frac{N_M + w}{N + k}. \text{ (From T4 and A3.)}$$

Consider a sequence of samples of increasing size N but such that $r = N_M/N$ remains constant. Then the value of $c^*(h_M, e_N)$ moves from w/k (for $N = 0$, i.e., tautological evidence) towards r (which is the limit for $N \rightarrow \infty$).

For further explanations and theorems on c^* see [Prob.] § 110.

XI. Further Axioms of Invariance (A8-A11)

A8. $c(h, e)$ is invariant with respect to any *permutation of the predicates* of any family.

T1. Let F be a family of k primitive predicates ' P_1 ', ..., ' P_k '. Let h_1, \dots, h_k be full sentences of these predicates with the same individual constant, and h be the disjunction of these sentences.

a. (Lemma.) For any e , $c(h, e) = 1$. (From A2, since h is L-true.)

b. Suppose that e' does not contain any predicate of F . Then for any i ($= 1, \dots, k$), $c(h_i, e') = 1/k$.

Proof. The k values $c(h_i, e')$ are equal (by A8). Their sum = $c(h, e')$ (by A3) = 1 (by (a)). Hence the assertion.

c. $m(h_i) = 1/k$. (From (b).)

A9. $c(h, e)$ is invariant with respect to any *permutation of families of the same size*.

- A10.** For non-general h and e , $c(h, e)$ is *independent of the total number of individuals*. (A10 corresponds to the requirement of a fitting c -sequence, [Prob.] § 57C.)
- A11.** $c(h, e)$ is *independent of the existence of other families* than those occurring in h or e .

XII. *Learning from experience* (A12)

The intuitive *principle of learning from experience* says that, other things being equal, the more frequently a kind of event has been observed, the more probable is its occurrence in the future. This is expressed more exactly in the *axiom of instantial relevance* (first proposed in Carnap [16])

- A12.** Suppose that e is non-L-false and non-general, and i and h are full sentences of the same factual, molecular predicate ' M ' with distinct individual constants not occurring in e .
- a. $c(h, e \cdot i) < c(h, e)$. ****THE '<' SYMBOL SHOULD HAVE A VERTICAL LINE THRU IT****
- b. $c(h, e \cdot i) \neq c(h, e)$.

Both c^\dagger (X) and the straight rule (VI) fulfill part (a) of A12, but violate part (b). With c^\dagger , i is always irrelevant for h . With the straight rule, i is irrelevant for h if e is a conjunction of full sentences of ' M '; in this case both c -values are 1.

- T1.** Let e , i , h , and M be as in A12.
- a. $c(h, e \cdot i) > c(h, e)$; i is positively relevant for h on e .
- b. Let j be a conjunction of n full sentences of ' M ' ($n \geq 2$) with n distinct individual constants which do not occur in e or h . Then
 $c(h, e \cdot j) > c(h, e)$. (From (a).)
- c. $c(h, e \cdot \sim i) < c(h, e)$; $\sim i$ is negatively relevant for h on e . (From (a) and [Prob.] T65-6e.)
- d. $c(h, e, i) > c(h, e \cdot \sim i)$. (From (a), (c).)

XIII. *The language L_F with one family F* (A13)

This and the subsequent sections refer to a language L_F whose primitive predicates are k predicates ' P_1 ', ..., ' P_k ' of a family F ($k \geq 2$). A sentence

in L_F may contain any number of individual constants but no variables. e_F is an individual distribution for s individuals with respect to F with the cardinal numbers s_i ($i = 1, \dots, k$). h_1, \dots, h_k are full sentences of ' P_1 ', ..., ' P_k ', respectively, with the same individual constant, which does not occur in e_F .

- A13.** Meaning postulates for F :
- a. $\vdash h_1 \vee h_2 \vee \dots \vee h_k$.
 - b. If $i \neq j$, $h_i \cdot h_j$ is L-false.

$m(e_F)$ is independent of other individuals (A10) and other families (A11). It depends not on the particular individuals in e_F but only on their numbers s_i . Therefore:

- T1.** For any m -function m fulfilling the axioms, there is, for any k , a representative mathematical function M_k of k arguments such that, for any e_F ,
- $$m(e_F) = M_k(s_1, s_2, \dots, s_k).$$

- T2.** M_k is invariant with respect to any permutation of the k arguments. (From A8.)

$e_F \cdot h_1$ is an individual distribution for $s + 1$ individuals with the cardinal numbers $s_1 + 1, s_2, \dots, s_k$. We define:

D1. $C_k(s_1; s_2, \dots, s_k) =_{\text{Df}} \frac{M_k(s_1 + 1, s_2, \dots, s_k)}{M_k(s_1, s_2, \dots, s_k)}$.

- T3.** a. For any c-function c and any k , there is a *representative mathematical function* C_k of k arguments such that, for any e_F , $c(h_i, e_F) = C_k(s_1; s_2, \dots, s_k)$. Analogously for h_2 , etc.
 b. C_k is invariant with respect to any permutation of the $k-1$ arguments following the first.

I shall sometimes write ' M ' and ' C ' without subscripts.

- T4.** For any k numbers n, p, s_3, \dots, s_k whose sum is s , the following holds. ('---' stands for ' s_3, \dots, s_k '; this expression drops out if $k = 2$; in this case $n + p = s$.)

$$\frac{C(n; p+1, ---)}{C(p; n+1, ---)} = \frac{C(n; p, ---)}{C(p; n, ---)}$$

(Here $p + n + 1 + \text{---} = s + 1$.)

Proof. The following holds identically:

$$\frac{M(n+1, p+1, \text{---})}{M(n, p+1, \text{---})} \times \frac{M(n, p+1, \text{---})}{M(n, p, \text{---})} = \frac{M(n+1, p+1, \text{---})}{M(n+1, p, \text{---})} \times \frac{M(n+1, p, \text{---})}{M(n, p, \text{---})}$$

According to D1, the first quotient is $C(n; p+1, \text{---})$; the second is (by T2) equal to

$$\frac{M(p+1, n, \text{---})}{M(p, n, \text{---})} = C(p; n, \text{---});$$

the third becomes (again with reordering of arguments) $C(p; n+1, \text{---})$, and the fourth $C(n; p, \text{---})$. Hence the theorem.

T5. a.
$$\sum_{i=1}^k c(h_i, e_F) = 1. \quad (\text{From A13a.})$$

b.
$$\sum_{i=1}^k C(s_i; s_1, \dots, s_{i-1}, s_{i+1}, \dots, s_k) = 1. \quad (\text{From (a).})$$

XIV. *The Axiom of Predictive Irrelevance* (A14)

Let e_1 be formed from e_F by replacing each predicate except ' P_1 ' with ' $\sim P_1$ '. Hence e_1 is an individual distribution for the s individuals with respect to the division $P_1, \sim P_1$, with the cardinal numbers s_1 and $s - s_1$. e_2, \dots, e_k are formed analogously.

For given k , $c(h_i, e_i)$ depends only on s_1 and s . It can therefore be represented by a function $G_k(s_1; s)$. Analogously for $i = 2, \dots, k$ (by A8).

T1. For any c -function c and any k , there is a *representative mathematical function* G_k such that, for $i = 1, \dots, k$.

$$c(h_i, e_i) = G_k(s_i; s).$$

T2. Suppose that $s_1 < s$. Let e'_1 be like e_1 but with the cardinal numbers $s_1 + 1$ and $s - s_1 - 1$.

a. $c(h_1, e'_1) > c(h_1, e_1)$. (From XII-T1d.)

b. $G_k(s_1 + 1; s) > G_k(s_1; s)$ (From (a).)

The axiom of predictive irrelevance says that of the k cardinal numbers in e_F all except s_1 are irrelevant for h_1

A14. For $k > 2$, $c(h_1, e_F) = c(h_1, e_1)$.

This axiom is not a necessary condition for the adequacy of c . But it is a customary (usually tacit) assumption, and it leads to a great simplification of the system. If $k = 2$, then e_1 is the same as e_F and therefore A14 is fulfilled trivially.

T3. For any $k (\geq 2)$ any any i :
a. $c(h_i, e_F) = c(h_i, e_i)$. (From A14, A8.)
b. For any numbers s_2, \dots, s_k whose sum is $s - s_1$, $C_k(s_1; s_2, \dots, s_k) = G_k(s_1; s)$. (From (a).)

I shall often write ‘ G ’ for ‘ G_k ’.

T4. For any sequence of k numbers s_1, \dots, s_k whose sum is s ,

$$\sum_{i=1}^k G(s_i; s) = 1. \text{ (From XIII-T5.)}$$

Special cases of T4:

T5. **a.** $G(s; s) + (k - 1) G(0; s) = 1$. (From T4 for the sequence $s, 0, \dots, 0$)
b. $G(s + 1; s + 1) = 1 - (k - 1) G(0; s + 1)$. (From (a).)
c. $G(1; 1) = 1 - (k - 1) G(0; 1)$. (From (a).)
d. $G(s; s + 1) + G(1; s + 1) + (k - 2) G(0; s + 1) = 1$. (Sequence $s, 1, 0, \dots, 0$.)

The following development has the aim to show (1) that, if all values of G for s are given, the values for $s + 1$ are uniquely determined, and (2) if $G(0; 1)$ is given, all values of G are uniquely determined. For these results it is *presupposed that $k > 2$* .

T6. For $k > 2$; for any n, p, s such that $n + p \leq s$.
a. $\frac{G(n; s + 1)}{G(p; s + 1)} = \frac{G(n; s)}{G(p; s)}$. (From XIII-T4.)
b. $G(n; s + 1) = G(0; s + 1) \frac{G(n; s)}{G(0; s)}$ (From (a) with $p = 0$.)

T7. For $k > 2$.

$$G(0; s + 1) \left[\frac{G(s; s)}{G(0; s)} + \frac{G(1; s)}{G(0; s)} + k - 2 \right] = 1.$$

(From T5d, by transforming the first two of its G -terms according to T 6b.)

Now aim (1) has been reached. If all G -values for s are given, $G(0; s + 1)$ is determined by T7, then the values $G(n; s + 1)$ for $n = 1, \dots, s$ are determined by T6b, and $G(s + 1; s + 1)$ by T5b. Thus all values for $s + 1$ are determined.

We have also attained aim (2). If $G(0; 1)$ is given, $G(1; 1)$ is determined by T5c. These are all the G -values for $s = 1$. They determine the values for $s = 2$, and so on. Thus all G -values are determined by $G(0; 1)$. The following theorem gives the explicit form.

T8. For $k > 2$, for any s and n ($0 \leq n \leq s$),

$$G(n; s) = \frac{n - (kn - 1)G(0; 1)}{s - (s - 1)kG(0; 1)}.$$

(This can be proved by mathematical induction with respect to s . (1) The theorem holds for $s = 1$ (for $n = 0$ it holds identically, for $n = 1$ by T5c). (2) If the theorem holds for a given s , it holds likewise for $s + 1$; this can be shown with the help of the theorems T7, T6b, and T5b, which determine the G -values for $s + 1$ on the basis of those for s . Hence the theorem holds for every s .)

Suppose that the value of $G(0; 1)$ has been chosen. Then all values of G can be determined. The following theorem T9 shows that the m -value of any state-description is determined by the value of G . Thereby the m -values for all sentences and the c -values for all pairs of sentences are determined (see VI).

T9. Let Z_F be a state-description for N individuals and for the k predicates of the family F , with the cardinal numbers N_i ($i = 1, \dots, k$). Then

$$m(Z_F) = \prod_i \prod_{n=0}^{N_i-1} G(n; S_i + n),$$

where \prod_i runs through those values of i for which $N_i > 0$;
 $S_i = \sum_{h=1}^{i-1} N_h$, $S_1 = 0$. (For the proof see [Continuum] § 5.)

XV. *The λ -system (A1 5)*

We shall construct a system of all c -functions fulfilling our axioms. We call it the λ -system, because the c -functions will be characterized by the values of a parameter λ .

We have seen that, for a given $k(> 2)$, all values of G are determined by $G(0; 1)$. The latter value can be freely chosen within certain boundaries. We shall now determine these boundaries.

From XI-T1c:

$$(1) \quad c(h_i, t) = 1/k.$$

Hence with XII-T1c (based on A12):

$$(2) \quad G(0; 1) < 1/k.$$

If we were to choose $G(0; 1) > 1/k$, then c would violate not only A12b but also A12a and therefore be unacceptable. If we choose $G(0; 1) = 1/k$, then only A12b is violated. This c does not belong to the λ -system, but will nevertheless be discussed as a boundary case (we shall find that it is the same as c^\dagger in X).

The following is obvious (from A1):

$$(3) \quad G(0; 1) \geq 0.$$

If we choose $G(0; 1) = 0$, then the resulting c fulfills A1 to A5, but violates A6. Hence it is quasi-regular. Therefore it does not belong to the λ -system. It will, however, be discussed as a boundary case; we shall see that it corresponds to the straight rule.

$$\mathbf{T1.} \quad 0 < G(0; 1) < 1/k.$$

It can also be shown that, if any value between 0 and $1/k$ is chosen for $G(0; 1)$, then the resulting c -function fulfills our axioms.

We define first an auxiliary parameter:

$$\mathbf{D1.} \quad \lambda_k''' =_{\text{Df}} kG_k(0; 1).$$

From T1:

$$\mathbf{T2.} \quad 0 < \lambda_k''' < 1.$$

I shall usually write ‘ λ''' ’ for ‘ λ_k''' ’. From XIV-T8:

T3. For $k > 2$, for any s and n ($0 \leq n \leq s$).

$$G(n; s) = \frac{n - (n - 1/k)\lambda'''}{s - (s - 1)\lambda'''}$$

We shall mostly use, not λ''' , but $\lambda = \lambda'''/(1 - \lambda''')$. The use of λ leads to a simpler formula for $G(n; s)$ (T4c). However, in the case of $G(0; 1) = 1/k$, $\lambda''' = 1$, while λ is infinite. Therefore in this case λ is less convenient than λ''' . But this case is not included in our system.

D2. $\lambda_k =_{\text{Df}} \frac{kG_k(0;1)}{1 - kG_k(0;1)}$.

T4. a. $\lambda''' = \lambda/(\lambda + 1)$.

b. $G(0; 1) = \frac{\lambda}{k(\lambda + 1)}$.

c. For $k > 2$, $G(n; s) = \frac{n + \lambda/k}{s + \lambda}$. (From T3, (a).)

The case $k = 2$. The important results at the end of XIV can be proved only if $k > 2$. (This seems surprising, since A14, on which the results are based, holds also for $k = 2$.) *A new axiom must be added for $k = 2$.* T4c shows that, for given $k(> 2)$, s , and $G(0; 1)$, $G(n; s)$ is a linear function of n . We assume as an axiom that the same holds for $k = 2$:

A15. For given s and $G_2(0; 1)$, $G(n; s)$ is a linear function of n .

Note that $G_2(n; s) = C_2(n; s - n)$ (see XVI-T3b). Therefore we have (without use of A15):

(4) $G_2(n; s) + G(s - n; s) = 1$. (From XIII-T5b.)

(5) $\frac{G_2(n; s + 1)}{G_2(s - n; s + 1)} = \frac{G_2(n; s)}{G_2(s - n; s)}$. (From XIII-T4.)

From (4) and (5) we can derive with A15:

T5. $G_2(n; s + 1) = G_2(0; s + 1) \frac{G_2(n; s)}{G_2(0; s)}$.

This corresponds to XIV-T6b. Then, in analogy to the earlier proofs, we can now prove the analogues of T3 and of T4c for $k = 2$; the latter is

$$\mathbf{T6.} \quad G_2(n; s) = \frac{n + \lambda / 2}{s + \lambda}.$$

For any λ , let c_λ be the c -function characterized by λ , and m_λ be the corresponding m -function. Using our result for $G(n; s)$ (T4c, T6), we obtain T7a from XIV-T9 (for the proofs of T7b and c and the notation $\left[\begin{smallmatrix} r \\ n \end{smallmatrix} \right]$, see [Continuum] § 10).

$$\begin{aligned} \mathbf{T7.} \quad \mathbf{a.} \quad m_\lambda(Z_F) &= \prod_{\substack{i \\ (N_i > 0)}} \prod_{n=0}^{N_i-1} \frac{n + \lambda / k}{S_i + n + \lambda} = \\ \mathbf{b.} \quad &= \frac{\prod_i \left[\left(\frac{\lambda}{k} \right) \left(\frac{\lambda}{k} + 1 \right) \left(\frac{\lambda}{k} + 2 \right) \dots \left(\frac{\lambda}{k} + N_i - 1 \right) \right]}{\lambda(\lambda + 1)(\lambda + 2) \dots (\lambda + N - 1)} \\ \mathbf{c.} \quad &= \frac{\prod_i \left[\begin{smallmatrix} N_i + \lambda / k - 1 \\ N_i \end{smallmatrix} \right]}{\left[\begin{smallmatrix} N_i + \lambda - 1 \\ N \end{smallmatrix} \right]}. \end{aligned}$$

d. Let Str_F be the structure-description corresponding to Z_F . Then

$$m_\lambda(Str_F) = \frac{\prod_i \left[\begin{smallmatrix} N_i + \lambda / k - 1 \\ N_i \end{smallmatrix} \right]}{\left[\begin{smallmatrix} N_i + \lambda - 1 \\ N \end{smallmatrix} \right]}.$$

(From (c) with VIII-T3.)

XVI. Various c -Functions in the λ -System

Let the predicate ' M ' be defined as a disjunction of w predicates of F : $P_1 \vee P_2 \vee \dots \vee P_w$. Hence its logical width is w . Let e_F be as before, and h_M a full sentence of ' M ' for a new individual. Let e_M be $s_1 + s_2 + \dots + s_w$. Then we have (from XV):

$$(1) \quad c_\lambda(h_M, e_F) = \frac{s_M + w\lambda / k}{s + \lambda}.$$

This is

$$\frac{\frac{s_M}{s} s + \frac{w}{k} \lambda}{s + \lambda},$$

thus it is the weighted mean of the observed relative frequency s_M/s and the relative width w/k , with weights s and λ , respectively. s_M/s is an empirical factor in the situation, and w/k is a logical factor. λ is thus the weight of the logical factor. The greater the chosen λ , the closer to w/k is the above c -value.

Example. For an even k , we take a predicate ‘ M ’ with $w = k/2$. In e_F , let $s = 10$, $s_M = 1$. Then the c_λ -value in (1), for various choices of λ , is as follows (see [Continuum] (12-19)):

λ	=	0	1	2	4	8	16	32	∞
$c_\lambda(h_M, e_F)$	=	0.1	0.136	0.167	0.214	0.278	0.346	0.405	0.5

For $\lambda = 0$, $c = s_M/s$. This is the straight rule, which violates A6. ([Continuum] § 14.)

For $\lambda = \infty$, $c = w/k = c(h_M, t)$. This is c^\dagger , which violates A 12b. ([Continuum] § 13.)

These are the two extreme methods, not included in the λ -system. In this system, we take $0 < \lambda < \infty$; hence the above c is between s_M/s and w/k (if these are unequal).

For families of different sizes (each in a separate language) we distinguish two kinds of inductive methods.

Inductive methods of the first kind: a fixed value is chosen for λ , independent of k . ([Continuum] § 11.)

Inductive methods of the second kind: λ_k is dependent upon k . The simplest form is: $\lambda_k = C_k$, with a constant C . The simplest method of this form takes $C = 1$, hence $\lambda_k = k$; thus from (1):

$$(2) \quad c_\lambda(h_M, e_F) = \frac{s_M + w}{s + k}.$$

This is the function c^ (see X).*

XVII. A Language with Two Families (A16)

The language \mathbb{L} contains two families: F^1 consists of k_1 predicates: ‘ P_1^1 ’, ‘ P_2^1 ’, etc; and F^2 of k_2 predicates: ‘ P_1^2 ’, ‘ P_2^2 ’, etc. There are

$k = k_1 k_2$ Q -predicates; Q_{ij} is the conjunction $P_i^1 \cdot P_j^2$ ($i = 1, \dots, k_1; j = 1, \dots, k_2$).

Let e^1 be an individual distribution for F^1 , and e^2 for F^2 , both for the same s individuals. Let e be $e^1 \cdot e^2$. This is an individual distribution for the k Q -predicates; let s_{ij} be the number of individuals with Q_{ij} .

We take the same λ for both families. Then we can determine $m_\lambda(e^1)$ and $m_\lambda(e^2)$ (by XV-T7).

Problem: What is to be taken as value of $m_\lambda(e)$? This is not determined by the previous axioms. We shall now consider two attempts at a solution, and then take a combination of them.

First tentative solution. We take the class of the k Q -predicates as the *pseudo-family* $F^{1,2}$. Then we define $m_\lambda^{1,2}$ for $F^{1,2}$, as if the latter were a real family; hence, in analogy to XV-T7c

$$\mathbf{D1.} \quad m_\lambda^{1,2}(e) = \text{Df} \frac{\prod_{i=1}^{k_1} \prod_{j=1}^{k_2} \left[\begin{array}{c} s_{ij} + \frac{\lambda}{k} - 1 \\ s_{ij} \end{array} \right]}{\left[\begin{array}{c} s + \lambda - 1 \\ s \end{array} \right]}.$$

$m_\lambda^{1,2}(e)$ depends only on the Q -numbers s_{ij} in e , not on the P -numbers in e^1 or e^2 .

Second tentative solution. We define $m^{\parallel 2}(e)$ as the product of the m -values for the two families separately:

$$\mathbf{D2.} \quad m_\lambda^{\parallel 2}(e) = \text{Df} = m_\lambda(e^1) \times m_\lambda(e^2).$$

$m^{\parallel 2}(e)$ depends only on the P -numbers, not on the Q -numbers.

We shall examine the two solutions with the help of the following three examples A, B, C , of individual distributions for $s = 20$ individuals, with $k_1 = k_2 = 2$ (the numerals in the four cells indicate the Q -numbers; the marginal numerals indicate the P -numbers for the two families).

		F^2	
		P_1^2	P_2^2
	10	5	5
F^1	P_1^1	10	5
	P_2^1	10	5

		F^2	
		P_1^2	P_2^2
	10	10	0
	10	0	10

		F^2	
		P_1^2	P_2^2
	20	10	10
	0	0	0

The following two *requirements* (or desiderata) I and II seem plausible.

(I) We should have: $m(A) < m(B)$, because B is more uniform than A .

This requirement is satisfied by $m^{1,2}$ (because the Q -numbers are equal in A , unequal in B), but not by $m^{||2}$ (this has equal values for A and for B , because the P -numbers are the same).

(II) We should have: $m(B) < m(C)$ because the distribution for F^1 is more uniform in C than in B , while that for F^2 is the same in C as in B .

This requirement is in accord with the customary analogy inference ('horse-donkey inference'). However, it is not satisfied by $m^{1,2}$ (this has equal values for B and for C , because the Q -numbers are the same). It is satisfied by $m^{||2}$.

Thus both solutions are unsatisfactory. Generally, any solution that uses only the P -numbers cannot satisfy I, and any solution that uses only the Q -numbers cannot satisfy II. An adequate solution must use both the P -numbers and the Q -numbers. This is done in the third solution, which satisfies both requirements.

Third solution. We define $m_{\lambda,n}(e)$ as a weighted mean of the first two solutions, with the weights η and $1-\eta$, where η is a new parameter

$$\mathbf{D3.} \quad m_{\lambda,n}(e) =_{\text{Df}} \eta m^{1,2}(e) + (1-\eta) m^{||2}(e).$$

The parameter η may be chosen, independently of λ , such that $0 < \eta < 1$. The greater η is, the stronger is the influence by analogy (i.e., the greater is the difference between the two c -values in A16 below). The method can easily be extended to more than two families; no new parameter is needed. (The method was worked out in collaboration with John Kemeny.)

The requirement II can be represented in a generalized form as follows:

A16. *Axiom of analogy.* Let e be an individual distribution for two families (with any k_1 and k_2). Let i and j be full sentences of Q_{11} and Q_{12} , respectively, with the same individual constant not occurring in e . Let h be a full sentence of Q_{12} with another individual constant not occurring in e . Then

$$c(h, e \cdot i) > c(h, e \cdot j).$$

The generalization for other Q -predicates follows by A8.

XVIII. *An Infinite Domain of Individuals* (A17)

Let the domain of L_N contain N individuals, and that of L_∞ be denumerably infinite. According to A10, the values of c for non-general sentences are in L_∞ the same as in L_N . If either e or h or both contain variables, a new axiom is needed. We take the value of c in L_∞ as the limit of its values in finite languages (see [Prob.] § 56)

A17. *Axiom of the infinite domain.* Let c_N be a c -function for L_N . Then the corresponding c -function c_∞ for L_∞ is determined as follows:

$$c_\infty(h, e) = \lim_{N \rightarrow \infty} c_N(h, e).$$

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